

Def: Let X, Y be rvs. The covariance of X and Y , write $\text{cov}(X, Y)$, is defined by

$$\text{cov}(X, Y) \equiv E[(X - \mu_X)(Y - \mu_Y)]$$

Prop: If X, Y are independent, then
 $\text{cov}(X, Y) = 0$

Prop: Let X_1, X_2, \dots, X_N ($1 \leq N < \infty$) be rvs.

Then

$$\text{var}(X_1 + \dots + X_N) = \sum_{i=1}^N \text{var}(X_i) + \sum_{i \neq j} \text{cov}(X_i, X_j)$$

In particular, if X_1, X_N are indep, then

$$\text{var}(X_1 + \dots + X_N) = \text{var}(X_1) + \dots + \text{var}(X_N)$$

$$\text{Pf: } \text{var}(X_1 + \dots + X_N) = E[(X_1 + \dots + X_N - \mu_{X_1 + \dots + X_N})^2]$$

$$= E\left[\left(\sum_{i=1}^N (X_i - \mu_{X_i})\right)^2\right]$$

$$= \sum_{i=1}^N E(X_i - \mu_{X_i})^2 + \sum_{i \neq j} E[(X_i - \mu_{X_i})(X_j - \mu_{X_j})]$$

□

Lemma: Let $a_1 < \dots < a_n$ and $(p_1, \dots, p_n) \in \mathbb{R}^n$
with $0 \leq p_i \leq 1$, $\sum_{i=1}^n p_i = 1$

Let $f = \{a_1, \dots, a_n\} \rightarrow \{p_1, \dots, p_n\}$ be a function
such that $f(a_i) = p_i$.

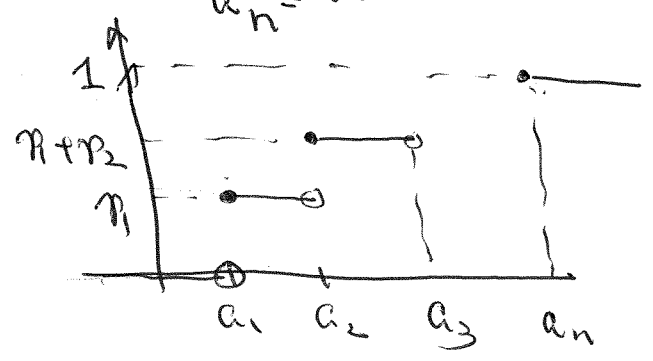
Then there is a prob space (Ω, \mathcal{F}, P)

and a rv $X: \Omega \rightarrow \{a_1, \dots, a_n\}$ such that

~~$f(a_i) = p_i$~~ $P\{X = a_i\} = p_i$ for all $i=1, \dots, n$.

pf: Define $F: \mathbb{R} \rightarrow \mathbb{R}$ by

$$F(t) = \begin{cases} 0 & t < a_1 \\ p_1 + \dots + p_i & a_i \leq t < a_{i+1} \\ 1 & a_n \leq t \end{cases}$$



Then $\exists (\Omega, \mathcal{F}, P)$ and a rv $X: \mathbb{R} \rightarrow \mathbb{R}$

st $F_X = F$ ie:

$$F_X(t) = F(t), \quad \forall t \in \mathbb{R}$$

Then $P\{X = a_i\} = P\{X \leq a_i\} - \lim_{h \rightarrow 0} P\{a_i - \frac{1}{h} < X \leq a_i\}$

$$= \lim_{h \rightarrow 0} [F_X(a_i) - F_X(a_i - \frac{1}{h})] = p_i$$

(39)

Def: Assume that there are n independent trials and the probability of success in each trial is p . Let X denote the number of successes. Then X is called a binomial random variable with parameter n, p . Write $X \sim b(n, p)$

Remark: There is a prob sp (Ω, \mathcal{E}, P) and a rv X st $X \sim b(n, p)$

Define: $\Omega = \{ \omega_1, \dots, \omega_n \mid x_i = 0 \text{ or } 1 \}$

$$\mathcal{E} = \mathcal{P}(\Omega)$$

Define $P\{ \omega_1, \dots, \omega_n \} = p^i (1-p)^{n-i}$

where $i =$ the number of x_k st $x_k = 1$

and

~~at~~ $X: \Omega \rightarrow \{0, 1, \dots, n\}$

$$X(\omega_1, \dots, \omega_n) = i$$

where $i =$ the number of x_k st $x_k = 1$.

Then $X \sim b(n, p)$

Prop: Let $X \sim b(n, p)$. Then

(i): $P\{X=k\} = \binom{n}{k} p^k (1-p)^{n-k}$, $0 \leq k \leq n$

(ii) $E(X) = np$

(iii) $Var(X) = np(1-p)$.

pf: (ii). $E(X) = \sum_{k=0}^n k P\{X=k\}$

$= \sum_{k=0}^n k \binom{n}{k} p^k (1-p)^{n-k} = \sum_{k=0}^n k \frac{n!}{k!(n-k)!} p^k (1-p)^{n-k}$

$= \sum_{k=1}^n \frac{n(n-1)!}{(k-1)!(n-k)!} p \cdot p^{k-1} (1-p)^{n-k}$

$= np (p + (1-p))^{n-1} = np$.

(iii). $var(X) = E(X^2) - E(X)^2$

$E(X^2) = \sum_{k=0}^n k^2 \frac{n!}{k!(n-k)!} p^k (1-p)^{n-k}$

$= \sum_{k=1}^n \frac{k n!}{(k-1)!(n-k)!} p^k (1-p)^{n-k}$

$= \sum_{k=1}^n \frac{(k-1) n!}{(k-1)!(n-k)!} p^k (1-p)^{n-k} + \sum_{k=1}^n \frac{n!}{(k-1)!(n-k)!} p^k (1-p)^{n-k}$

$= \sum_{k=2}^n \frac{n!}{(k-2)!(n-k)!} p^k (1-p)^{n-k} + \sum_{j=0}^{n-1} \frac{n!}{j!(n-1-j)!} p^{j+1} (1-p)^{n-1-j}$

$= \sum_{i=0}^{n-2} \frac{n!}{i!(n-2-i)!} p^{i+2} (1-p)^{n-2-i} +$

$$= p^2 n(n-1) [p + (1-p)]^{n-2} + np (p + (1-p))^{n-1}$$

(4)

$$\therefore \text{Var}(X) = E(X^2) - E(X)^2$$

$$= n(n-1)p^2 + np - n^2p^2$$

$$= -np^2 + np = np(1-p) \quad \square$$

Alternative proof =

Define $X_i: \Omega \rightarrow \mathbb{R}$ by

$$X_i = \begin{cases} 1 & \text{if } i\text{-th trial success} \\ 0 & \text{otherwise} \end{cases}$$

Then $X = X_1 + \dots + X_n$

Note: $E(X_i) = p$

$$\begin{aligned} \text{Var}(X_i) &= E(X_i^2) - E(X_i)^2 \\ &= E(X_i) - E(X_i)^2 \\ &= p - p^2 \end{aligned}$$

∴

□

Observation:

(42)

let $X_n \sim \text{bin}(n, p)$ and $\lambda = np$.

Fix λ and $k \in \mathbb{N}$.

$$\begin{aligned} P\{X_n = k\} &= \binom{n}{k} p^k (1-p)^{n-k} \\ &= \frac{n(n-1)\dots(n-k+1)}{k!} \cdot \left(\frac{\lambda}{n}\right)^k \left(1 - \frac{\lambda}{n}\right)^{n-k} \\ &= 1 \cdot \left(1 - \frac{\lambda}{n}\right) \dots \left(1 - \frac{k-1}{n}\right) \cdot \frac{1}{k!} \lambda^k \left(1 - \frac{\lambda}{n}\right)^n \left(1 - \frac{\lambda}{n}\right)^{-k} \\ &\rightarrow \frac{1}{k!} \lambda^k e^{-\lambda} \end{aligned}$$

Def: A discrete rv $X: \Omega \rightarrow \{0, 1, 2, \dots\}$ is called a Poisson rv with parameter $\lambda > 0$ if

$$P\{X = k\} = \frac{1}{k!} e^{-\lambda} \lambda^k, \quad \forall k = 0, 1, 2, \dots$$

Remark: Define $F: \mathbb{R} \rightarrow \mathbb{R}$ by

$$F(t) = \begin{cases} 0 & t < 0 \\ \frac{1}{k!} e^{-\lambda} \lambda^k & k \leq t < k+1 \\ \sum_{i=0}^k \frac{1}{i!} e^{-\lambda} \lambda^i & k \leq t < k+1 \end{cases}$$

Note F is right cont and

(43)

$$\lim_{t \rightarrow \infty} F(t) = \lim_{k \rightarrow \infty} \sum_{i=0}^k e^{-\lambda} \frac{\lambda^i}{i!} = 1$$

\therefore Such a Poisson rv exists ~~forall~~. \square

Prop: If X is a Poisson rv with a parameter $\lambda > 0$,
then $E(X) = \lambda = \text{var}(X)$

pf: 1: $E(X) = \lambda$

$$\text{Note: } E(X) = \sum_{k=0}^{\infty} k P\{X=k\} = \sum_{k=0}^{\infty} k \cdot \frac{1}{k!} \lambda^k e^{-\lambda}$$

$$= \left(\sum_{k=1}^{\infty} \frac{1}{(k-1)!} \lambda^k \right) e^{-\lambda}$$

$$= \left(\sum_{j=0}^{\infty} \frac{1}{j!} \lambda^{j+1} \right) e^{-\lambda} = \lambda$$

$(j = k-1)$

2: $\text{var}(X) = \lambda$: Note $\text{var}(X) = E(X^2) - E(X)^2$

$$\text{Note: } E(X^2) = \sum_{k=0}^{\infty} k^2 P\{X=k\} = \sum_{k=1}^{\infty} k^2 e^{-\lambda} \frac{\lambda^k}{k!}$$

$$= \sum_{k=2}^{\infty} e^{-\lambda} \frac{\lambda^k}{(k-2)!} + \left(\sum_{k=1}^{\infty} \frac{\lambda^k}{(k-1)!} \right) e^{-\lambda}$$

$$= e^{-\lambda} \sum_{k=2}^{\infty} \frac{\lambda^k}{(k-2)!} + \left(\sum_{j=0}^{\infty} \frac{\lambda^{j+1}}{j!} \right) e^{-\lambda}$$

$$= e^{-\lambda} \left(\sum_{i=0}^{\infty} \frac{\lambda^i}{i!} \right) \lambda^2 + \left(\sum_{j=0}^{\infty} \frac{\lambda^j}{j!} \right) \lambda e^{-\lambda}$$

$(j = k-1)$

$$= \lambda^2 + \lambda \quad \therefore \text{var}(X) = \lambda \quad \square$$